## ПAmIBIA UПIVERSITY

OF SCIEПCE AПD TECHחOLOGY

## FACULTY OF HEALTH, APPLIED SCIENCES AND NATURAL RESOURCES

DEPARTMENT OF MATHEMATICS AND STATISTICS

| QUALIFICATION: Bachelor of Science Honours in Applied Statistics |  |
| :--- | :--- |
| QUALIFICATION CODE: 08BSHS | LEVEL: 8 |
| COURSE CODE: MVA802S | COURSE NAME: MULTIVARIATE ANALYSIS |
| SESSION: NOVEMBER 2022 | PAPER: THEORY |
| DURATION: 3 HOURS | MARKS: 100 |


| FIRST OPPORTUNITY EXAMINATION QUESTION PAPER |  |
| :--- | :---: |
| EXAMINER | Dr D. B. GEMECHU |
|  |  |
| MODERATOR: | Prof L. PAZVAKAWAMBWA |

## INSTRUCTIONS

1. There are 8 questions, answer ALL the questions by showing all the necessary steps.
2. Write clearly and neatly.
3. Number the answers clearly.
4. Round your answers to at least four decimal places, if applicable.

PERMISSIBLE MATERIALS

1. Non-programmable scientific calculator

THIS QUESTION PAPER CONSISTS OF 6 PAGES (Including this front page)
ATTACHMENTS
Two statistical distribution tables (z-and F-distribution tables)

## Question 1 [11 Marks]

1.1. Briefly discuss a one-way MANOVA. Your answer should include (Definition, three assumptions of one-way MANOVA, hypothesis to be tested under one-way MANOVA and two of the most common test statistics used to test the hypothesis).
[1+2+2]
1.2. Briefly discuss two-sample profile analysis. Your answer should include the definition, the assumptions, and the possible hypothesis of interests that can be tested using this approach.
$[1+2+3]$

## Question 2 [9 Marks]

2. The data in table below are three measurements on air-pollution variables recorded on three different days.

| Days | Solar <br> radiation, <br> $y_{1}$ | Nitrogen <br> Dioxide <br> $\left(\mathrm{NO}_{2}\right)$, <br> $y_{2}$ | Ozone <br> $\left(\mathrm{O}_{3}\right), y_{3}$ |
| :---: | :---: | :---: | :---: |
| 1 | 72 | 18 | 9 |
| 2 | 70 | 11 | 7 |
| 3 | 80 | 13 | 11 |

Assume that $y \sim N_{3}(\mu, \Sigma)$ with unknown $\mu$ and unknown $\Sigma$. Then, using the matrices approach, calculate the maximum likelihood estimate of the population:
2.1. mean vector.
2.2. variance-covariance matrix.

## Question 3 [10 Marks]

3. If $\boldsymbol{y} \sim N_{p}(\boldsymbol{\mu}, \boldsymbol{\Sigma})$ and $\boldsymbol{z}=\left(\boldsymbol{\Sigma}^{\mathbf{1 / 2}}\right)^{-1}(\boldsymbol{y}-\boldsymbol{\mu})$, then show that $\boldsymbol{z} \sim N_{p}(\mathbf{0}, \boldsymbol{I})$. Hint: Use the uniqueness property of joint moment generating function.

## Question 4 [11 marks]

4. Perspiration from 19 healthy females was analyzed. Two components, $y_{1}=$ sweat rate, and $y_{2}=$ sodium, were measured. Assume that the data is from a multivariate normal distribution $N_{2}(\boldsymbol{\mu}, \boldsymbol{\Sigma})$ with unknown $\boldsymbol{\mu}$ and unknown $\boldsymbol{\Sigma}$. The mean score and covariance matrix of the score are:

$$
\begin{gathered}
\overline{\boldsymbol{y}}=\binom{4.640}{45.400} \\
\boldsymbol{S}=\left(\begin{array}{cc}
2.879 & 10.010 \\
10.010 & 199.788
\end{array}\right)
\end{gathered}
$$

Test the hypothesis $H_{0}: \boldsymbol{\mu}=(4,50)^{\prime}$ vs $H_{1}: \boldsymbol{\mu} \neq(4,50)^{\prime}$ at $5 \%$ level of significance. Your solution should include the following:
4.1. State the test statistics to be used and its corresponding distribution
4.2. State the decision (rejection) rule and compute the tabulated value using an appropriate statistical table
4.3. Compute the test statistics and write up your decision and conclusion
5. Two psychological tests were given to 11 men and 10 women. The variables are $y_{1}=$ tool recognition and $y_{2}=$ vocabulary. The mean vectors and covariance matrices of the two samples are $\quad \bar{y}_{1}=\binom{12}{13}, \bar{y}_{2}=\binom{16}{17}, S_{1}=\left(\begin{array}{cc}5 & 4 \\ 4 & 13\end{array}\right)$ and $S_{2}=\left(\begin{array}{cc}9 & 7 \\ 7 & 18\end{array}\right)$.
Assume that the observations are bivariate and follow multivariate normal distributions $N\left(\boldsymbol{\mu}_{i}, \mathbf{\Sigma}\right)$, for $i=1$ and 2 .
5.1. Compute the pooled covariance matrix
5.2. Conduct a test if there is any significant difference between the vector of expected mean scores of men and women at $5 \%$ level of significance. Your answer should include the following:
5.2.1. State the null and alternative hypothesis to be tested
5.2.2. State the test statistics to be used and its corresponding distribution
5.2.3. State the decision (rejection) rule and compute the tabulated value using an appropriate statistical table
[3]
5.2.4. Compute the test statistics and write up your decision and conclusion

## Question 6 [23 Marks]

6. Let $\boldsymbol{x} \sim N_{5}(\boldsymbol{\mu}, \boldsymbol{\Sigma})$, where $\boldsymbol{x}=\left(\begin{array}{l}x_{1} \\ x_{2} \\ x_{3} \\ x_{4} \\ x_{5}\end{array}\right), \boldsymbol{\mu}=\left(\begin{array}{l}5 \\ 3 \\ 7 \\ 4 \\ 9\end{array}\right) \quad$ and $\quad \boldsymbol{\Sigma}=\left(\begin{array}{ccccc}4 & -1 & 0 & 0 & 2 \\ -1 & 4 & 2 & 0 & 4 \\ 0 & 2 & 9 & 0 & 3 \\ 0 & 0 & 0 & 9 & 7 \\ 2 & 4 & 3 & 7 & 16\end{array}\right)$.

Answer the following questions based on the above information.
6.1. If $z_{1}=\frac{x_{1}+x_{3}}{2}$ and $z_{2}=x_{1}-\frac{1}{2} x_{2}$ then, find the joint distribution of $z_{1}$ and $z_{2}$. Are they independently distributed? Provide explanation for your answer.
6.2. Find the conditional distribution of $x_{2}$ given $\left(x_{1}, x_{3}\right)$.
6.3. If $y=2 x_{1}-3 x_{2}+x_{3}$, then find $P(y>7)$

## Question 7 [9 Marks]

7. Let $\boldsymbol{X}^{\prime}=\left[X_{1}, X_{2}, \ldots, X_{p}\right]$ have covariance matrix $\boldsymbol{\Sigma}$ with eigenvalue-eigenvector pairs $\left(\lambda_{1}, e_{1}\right),\left(\lambda_{2}, e_{2}\right), \ldots,\left(\lambda_{p}, e_{p}\right)$ where $\lambda_{1} \geq \lambda_{2} \geq \cdots \geq \lambda_{p} \geq 0$. Let $Y_{i}=e_{i}^{\prime} X, Y_{2}=e_{2}^{\prime} X, \ldots, Y_{p}=$ $\boldsymbol{e}_{p}^{\prime} \boldsymbol{X}$ be the principal components. Then show that
7.1. $\operatorname{Var}\left(Y_{i}\right)=\lambda_{i}$
7.2. $\operatorname{tr}(\boldsymbol{\Sigma})=\sum_{i=1}^{p} \operatorname{Var}\left(Y_{i}\right)=\lambda_{1}+\lambda_{2}+\cdots+\lambda_{p}$

## Question 8 [13 marks]

8. A researcher compared judges' scores on fish prepared by three methods. Twelve fish were cooked by each method, and several judges tasted fish samples and rated each on four variables: $y_{1}=$ aroma, $y_{2}=$ flavor, $y_{3}=$ texture, and $y_{4}=$ moisture. The summary statistics of the data are given in the attached software output (Tables 1-5 given below).
8.1. Draw conclusion of the Box test for equality of covariance matrix using the $5 \%$ significance level. Your answer should include the hypothesis to be tested, test statics and $p-v a l u e$ and conclusion.
8.2. Are there significant mean difference of judges' scores (as rated each on four variables) between three different methods? Your answer should include the hypothesis to be tested, test statics and $p-v a l u e$ and conclusion.
8.3. Are there significant mean difference of judges' score on flavour of fish prepared by three methods? If so, which cooking methods differ?
8.4. Are there significant mean difference judges' score on moisture of fish prepared by three methods? Explain in detail.

Table 1: Box's Test of Equality of Covariance Matrices ${ }^{\text {a }}$

| Box's M | 16.292 |
| :--- | :--- |
| $F$ | .669 |
| df1 | 20 |
| df2 | 3909.028 |
| Sig. | .860 |

Tests the null hypothesis that the observed covariance matrices of the dependent variables are equal across groups.
a. Design: Intercept + Method

Table 2: Multivariate Tests ${ }^{\text {a }}$

| Effect |  | Value | F | Hypoth esis df | Error df | Sig. | Partial Eta Squared |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Intercept | Pillai's Trace | . 993 | $1109.613^{\text {b }}$ | 4.000 | 30.000 | . 000 | . 993 |
|  | Wilks' Lambda | . 007 | $1109.613^{\text {b }}$ | 4.000 | 30.000 | . 000 | . 993 |
|  | Hotelling's Trace | 147.948 | $1109.613^{\text {b }}$ | 4.000 | 30.000 | . 000 | . 993 |
|  | Roy's Largest Root | 147.948 | $1109.613^{\text {b }}$ | 4.000 | 30.000 | . 000 | . 993 |
| Method | Pillai's Trace | . 864 | 5.897 | 8.000 | 62.000 | . 000 | . 432 |
|  | Wilks' Lambda | . 220 | $8.488^{\text {b }}$ | 8.000 | 60.000 | . 000 | . 531 |
|  | Hotelling's Trace | 3.162 | 11.461 | 8.000 | 58.000 | . 000 | . 613 |
|  | Roy's Largest Root | 3.036 | $23.526^{\text {c }}$ | 4.000 | 31.000 | . 000 | . 752 |

a. Design: Intercept + Method
b. Exact statistic
c. The statistic is an upper bound on $F$ that yields a lower bound on the significance level.

Table 3: Levene's Test of Equality of Error Variances ${ }^{\text {a }}$

|  |  | Levene <br> Statistic | df1 | df2 | Sig. |
| :--- | :--- | ---: | ---: | ---: | ---: |
| flavor | Based on Mean | .158 | 2 | 33 | .855 |
|  | Based on Median | .245 | 2 | 33 | .784 |
|  | Based on Median and <br> with adjusted df | .245 | 2 | 32.566 | .784 |
|  | Based on trimmed mean | .166 | 2 | 33 | .848 |
|  | Based on Mean | .592 | 2 | 33 | .559 |
|  | Based on Median | .547 | 2 | 33 | .584 |
|  | Based on Median and <br> with adjusted df | .547 | 2 | 32.090 | .584 |
|  | Based on trimmed mean | .588 | 2 | 33 | .561 |
| moisture | Based on Mean | 1.167 | 2 | 33 | .324 |
|  | Based on Median | 1.263 | 2 | 33 | .296 |
|  | Based on Median and | 1.263 | 2 | 32.455 | .296 |
|  | with adjusted df |  |  |  |  |


|  | Based on trimmed mean | 1.195 | 2 | 33 | .316 |
| :--- | :--- | ---: | ---: | ---: | ---: |
| aroma | Based on Mean | .684 | 2 | 33 | .512 |
|  | Based on Median | .680 | 2 | 33 | .514 |
|  | Based on Median and <br> with adjusted df | .680 | 2 | 31.390 | .514 |
|  | Based on trimmed mean | .695 | 2 | 33 | .506 |

Tests the null hypothesis that the error variance of the dependent variable is equal across groups.
a. Design: Intercept + Method

Table 4: Tests of Between-Subjects Effects

| Source | Dependent Variable | Type III Sum of Squares | df | Mean <br> Square | F | Sig. | Partial Eta Squared |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Corrected Model | Flavour | $4.605^{\text {a }}$ | 2 | 2.303 | 9.378 | . 001 | . 362 |
|  | Texture | $2.382^{\text {b }}$ | 2 | 1.191 | 3.386 | . 046 | . 170 |
|  | Moisture | .811 ${ }^{\text {c }}$ | 2 | . 405 | 1.266 | . 295 | . 071 |
|  | Aroma | $1.051^{\text {d }}$ | 2 | . 525 | 1.293 | . 288 | . 073 |
| Intercept | Flavour | 995.402 | 1 | 995.402 | 4054.092 | . 000 | . 992 |
|  | Texture | 1110.000 | 1 | 1110.000 | 3155.719 | . 000 | . 990 |
|  | Moisture | 1309.234 | 1 | 1309.234 | 4089.096 | . 000 | . 992 |
|  | Aroma | 975.521 | 1 | 975.521 | 2400.910 | . 000 | . 986 |
| Method | Flavour | 4.605 | 2 | 2.303 | 9.378 | . 001 | . 362 |
|  | Texture | 2.382 | 2 | 1.191 | 3.386 | . 046 | . 170 |
|  | Moisture | . 811 | 2 | . 405 | 1.266 | . 295 | . 071 |
|  | Aroma | 1.051 | 2 | . 525 | 1.293 | . 288 | . 073 |
| Error | Flavour | 8.103 | 33 | . 246 |  |  |  |
|  | Texture | 11.607 | 33 | . 352 |  |  |  |
|  | Moisture | 10.566 | 33 | . 320 |  |  |  |
|  | Aroma | 13.408 | 33 | . 406 |  |  |  |
| Total | Flavour | 1008.110 | 36 |  |  |  |  |
|  | Texture | 1123.990 | 36 |  |  |  |  |
|  | Moisture | 1320.610 | 36 |  |  |  |  |
|  | Aroma | 989.980 | 36 |  |  |  |  |
| Corrected Total | Flavour | 12.708 | 35 |  |  |  |  |
|  | Texture | 13.990 | 35 |  |  |  |  |
|  | Moisture | 11.376 | 35 |  |  |  |  |
|  | Aroma | 14.459 | 35 |  |  |  |  |

a. R Squared $=.362$ (Adjusted R Squared $=.324$ )
b. R Squared $=.170$ (Adjusted R Squared $=.120$ )
c. R Squared $=.071$ (Adjusted R Squared $=.015$ )
d. $R$ Squared $=.073$ (Adjusted R Squared $=.016$ )

Table 5: Pairwise Comparisons

| Dependent <br> Variable | (I) Method | (J) Method | Mean <br> Differe <br> nce (I-J) | Std. <br> Error | Sig. ${ }^{\text {b }}$ | 95\% Confidence Interval for Difference ${ }^{\text {b }}$ |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  |  |  | Lower Bound | Upper Bound |
| flavour | 1 | 2 | . 475 | . 202 | . 075 | -. 035 | . 985 |
|  |  | 3 | .875* | . 202 | . 000 | . 365 | 1.385 |
|  | 2 | 1 | -. 475 | . 202 | . 075 | -. 985 | . 035 |
|  |  | 3 | . 400 | . 202 | . 169 | -. 110 | . 910 |
|  | 3 | 1 | -.875* | . 202 | . 000 | -1.385 | -. 365 |
|  |  | 2 | -. 400 | . 202 | . 169 | -. 910 | . 110 |
| texture | 1 | 2 | . 133 | . 242 | 1.000 | -. 477 | . 744 |
|  |  | 3 | -. 467 | . 242 | . 188 | -1.077 | . 144 |
|  | 2 | 1 | -. 133 | . 242 | 1.000 | -. 744 | . 477 |
|  |  | 3 | -. 600 | . 242 | . 055 | -1.211 | . 011 |
|  | 3 | 1 | . 467 | . 242 | . 188 | -. 144 | 1.077 |
|  |  | 2 | . 600 | . 242 | . 055 | -. 011 | 1.211 |
| moisture | 1 | 2 | . 108 | . 231 | 1.000 | -. 474 | . 691 |
|  |  | 3 | -. 250 | . 231 | . 861 | -. 833 | . 333 |
|  | 2 | 1 | -. 108 | . 231 | 1.000 | -. 691 | . 474 |
|  |  | 3 | -. 358 | . 231 | . 391 | -. 941 | . 224 |
|  | 3 | 1 | . 250 | . 231 | . 861 | -. 333 | . 833 |
|  |  | 2 | . 358 | . 231 | . 391 | -. 224 | . 941 |
| aroma | 1 | 2 | . 125 | . 260 | 1.000 | -. 531 | . 781 |
|  |  | 3 | . 408 | . 260 | . 378 | -. 248 | 1.065 |
|  | 2 | 1 | -. 125 | . 260 | 1.000 | -. 781 | . 531 |
|  |  | 3 | . 283 | . 260 | . 852 | -. 373 | . 940 |
|  | 3 | 1 | -. 408 | . 260 | . 378 | -1.065 | . 248 |
|  |  | 3 | -. 283 | . 260 | . 852 | -. 940 | . 373 |

Based on estimated marginal means
*. The mean difference is significant at the .05 level.
b. Adjustment for multiple comparisons: Bonferroni.

```
=== END OF PAPER===
    TOTAL MARKS: }10
```



|  | 0.00 | 0.01 | 0.02 | 0.03 | 0.0 | 0.05 | 0.06 | 0.07 | 0.08 | 0.09 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 0. | 0.0000 | 0.0040 | 0.0080 | 0.0120 | 0.0160 | 0.0199 | 0.0239 | 0.0279 | 0.0319 | 0.0359 |
| 0 | 0.0398 | 0.0438 | 0.0478 | 0.0517 | 0.0557 | 0.0596 | 0.0636 | 0.0675 | 0.0714 | 0.0753 |
| 0.2 | 0.079 | 0.0832 | 0.0 | 0.0 | 0. |  | 0.1 | 0.1064 | 0.1103 | 0.1141 |
| 0.3 | 0.1 | 0.1217 | 0.125 | 0 | 0. | 0. | 0. | 0.1443 | 0.1480 | . 1517 |
| 0.4 | 0.1 | 0.1 | 0.16 | 0.1664 | 0.1 | 0. | 0.17 | 0.1808 | 0.1844 | 0.1879 |
| 0 | 0.1 | 0.1 | 0.1 | 0 | 0. | 0.2088 | 0.2123 | 0.2157 | 0.2190 | 0.2224 |
| 0. | 0.225 | 0.22 | 0.232 | 0.235 | 0.2 | 0.2422 | 0.2 | 0.2486 | 0.2517 | 0.2549 |
| 0.7 | 0.25 |  | 0.2 | 0 |  | 0.2 | 0.2 | 0.2794 | 0.2823 | 0.2852 |
| . | 0. | 0.2 | 0.2 | 0.2967 | 0. | 0.30 | 0.3051 | 0.3078 | 0.3106 | 0.3133 |
| 0 | 0.3 | 0.3186 | 0.3 | 0.3238 | 0.3 | 0. | 0.3315 | 0.3340 | 0.3365 | 0.3389 |
| 1. | 0.341 | 0.3438 | 0.346 | 0.3485 | 0.350 | 0.353 | 0.35 | 0.3577 | 0.3599 | 0.3621 |
| 1 | 0.364 | 0.366 | 0.368 | 0.3708 |  | 0.37 | 0.3 | 0.3790 | 0.3810 | 0.3830 |
| 1 | 0. | 0.38 | 0.388 | 0.3907 | 0.3 | 0.39 | 0.39 | 0.3980 | 0.3997 | 0.4015 |
| 1 | 0. | 0 | 0. | 0.4082 | 0 | 0. | 0.4131 | 0.4147 | 0.4162 | 7 |
| 1. | 0.419 | 420 | 0.422 | 0.4236 | 0.425 | 0.4265 | 0.42 | 0.4292 | 0.4306 | 0.4319 |
| 1. | 0.433 | 0.43 | 0. | 0.4370 |  | 0.4 | 6 | 0.4418 | 0.4429 | 0.4441 |
| 1. | 0.445 | 0.446 | 0. |  | 0. | 0.45 | . 45 | 0.4525 | 0.4535 | 0.4545 |
| 1.7 | 0.455 | 0. | 0.457 | 0.45 | 0.4 | 0.4599 | 0.4608 | 0.4616 | 0.4625 | 0.4633 |
| 1.8 | 0.46 | 0.46 | 0.465 | 0 | 0.4 | 0.4 | 0.4686 | 0.4693 | 0.4699 | 0.4706 |
| 1.9 | 0.471 | 0. | 0.472 | 0 |  | 0.4 | 0.4750 | 0.4756 | 0.4761 | 0.4767 |
| 2.0 | 0.477 | 0.4778 | 0.478 | 0.4 | 0.4 | 0.4798 | 0.4803 | 0.4808 | 0.4812 | 0.4817 |
| 2. | 0.482 | 0.4826 | 0.483 | 0.48 | 0.483 | 0.4842 | 0.4846 | 0.4850 | 0.4854 | 0.4857 |
| 2.2 | 0.486 | 0.48 | 0.486 | 0 | 0.48 | 0.4878 | 0.48 | 0.4884 | 0.4887 | 0.4890 |
| 2.3 | 0.489 | 0.48 | 0.4 | 0. | 0 | 0.4906 | 0.4909 | 0.4911 | 0.4913 | 0.4916 |
| 2.4 | 0.491 | 0.4920 | 0.4922 | 0.4 | 0. | 0.4929 | 0.4931 | 0.4932 | 0.4934 | 0.4936 |
| 2.5 | 0.493 | 0.4940 | 0.4941 | 0.4943 | 0.4945 | 0.4946 | 0.4948 | 0.4949 | 0.4951 | 0.4952 |
| 2.6 | 0.4953 | 0.4955 | 0.495 | 0.495 | 0.495 | 0.4960 | 0.49 | 0.4962 | 0.4963 | 0.4964 |
| 2.7 | 0.4965 | 0.4966 | 0.496 | 0.4968 | 0.4969 | 0.4970 | 0. | 0.4972 | 0.4973 | 0.4974 |
| 2.8 | 0.4974 | 0.4975 | 0.4976 | 0.4977 | 0.4977 | 0.4978 | 0.4979 | 0.4979 | 0.4980 | 0.4981 |
| 2.9 | 0.4981 | 0.4982 | 0.4982 | 0.4983 | 0.4984 | 0.4984 | 0.4985 | 0.4985 | 0.4986 | 0.4986 |
| 3.0 | 0.4987 | 0.4987 | 0.4987 | 0.4988 | 0.4988 | 0.4989 | 0.4989 | 0.4989 | 0.4990 | 0.4990 |

Table for $\alpha=.05$


| df2/df1 | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10 | 12 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1 | 161.448 | 199.500 | 215.707 | 224.583 | 230.162 | 233.986 | 236.768 | 238.883 | 240.543 | 241.882 | 243.906 |
| 2 | 18.513 | 19.000 | 19.164 | 19.247 | 19.296 | 19.329 | 19.353 | 19.371 | 19.384 | 19.396 | 19.413 |
| 3 | 10.128 | 9.552 | 9.277 | 9.117 | 9.014 | 8.941 | 8.887 | 8.845 | 8.812 | 8.786 | 8.745 |
| 4 | 7.709 | 6.944 | 6.591 | 6.388 | 6.256 | 6.163 | 6.0942 | 6.041 | 5.998 | 5.964 | 5.912 |
| 5 | 6.608 | 5.786 | 5.409 | 5.192 | 5.050 | 4.950 | 4.876 | 4.818 | 4.772 | 4.735 | 4.678 |
| 6 | 5.987 | 5.143 | 4.757 | 4.533 | 4.387 | 4.284 | 4.207 | 4.147 | 4.099 | 4.060 | 3.999 |
| 7 | 5.591 | 4.737 | 4.347 | 4.120 | 3.972 | 3.866 | 3.787 | 3.726 | 3.676 | 3.637 | 3.575 |
| 8 | 5.318 | 4.459 | 4.066 | 3.838 | 3.688 | 3.581 | 3.501 | 3.438 | 3.388 | 3.347 | 3.284 |
| 9 | 5.117 | 4.256 | 3.863 | 3.633 | 3.482 | 3.374 | 3.293 | 3.229 | 3.178 | 3.137 | 3.073 |
| 10 | 4.965 | 4.103 | 3.708 | 3.478 | 3.326 | 3.217 | 3.136 | 3.072 | 3.020 | 2.978 | 2.913 |
| 11 | 4.844 | 3.982 | 3.587 | 3.358 | 3.204 | 3.095 | 3.012 | 2.948 | 2.896 | 2.854 | 2.788 |
| 12 | 4.747 | 3.885 | 3.490 | 3.259 | 3.106 | 2.996 | 2.913 | 2.849 | 2.796 | 2.753 | 2.687 |
| 13 | 4.667 | 3.806 | 3.411 | 3.179 | 3.025 | 2.915 | 2.832 | 2.767 | 2.714 | 2.671 | 2.604 |
| 14 | 4.600 | 3.739 | 3.344 | 3.112 | 2.958 | 2.848 | 2.764 | 2.699 | 2.645 | 2.602 | 2.534 |
| 15 | 4.543 | 3.682 | 3.287 | 3.056 | 2.901 | 2.791 | 2.707 | 2.641 | 2.587 | 2.544 | 2.475 |
| 16 | 4.494 | 3.634 | 3.239 | 3.007 | 2.852 | 2.741 | 2.657 | 2.591 | 2.537 | 2.494 | 2.425 |
| 17 | 4.451 | 3.591 | 3.197 | 2.965 | 2.810 | 2.699 | 2.614 | 2.548 | 2.494 | 2.450 | 2.381 |
| 18 | 4.414 | 3.555 | 3.160 | 2.928 | 2.773 | 2.661 | 2.577 | 2.510 | 2.456 | 2.412 | 2.342 |
| 19 | 4.381 | 3.522 | 3.127 | 2.895 | 2.740 | 2.628 | 2.544 | 2.477 | 2.423 | 2.378 | 2.308 |
| 20 | 4.351 | 3.493 | 3.098 | 2.866 | 2.711 | 2.599 | 2.514 | 2.441 | 2.393 | 2.348 | 2.278 |

